

Spatial Statistics

GEOG 419: Lembo

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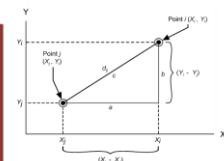
Point Pattern Analysis

- Global methods to analyze point patterns across entire study region (or a map)
 - Quantitative tools for examining a spatial arrangement of point locations on the landscape
- Two common types of analysis
 - spacing of individual points – nearest neighbor analysis
 - Ex. fire stations locations – random or dispersed
 - Goal: equitable service throughout region
 - Design new configuration (e.g., relocating, new stations)
 - More or less dispersed than original configuration
 - nature of overall point pattern – are locations dispersed or clustered
 - Ex. diseased trees in a national forest
 - Widespread aerial spraying versus concentrated ground treatment

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Center Point



Pythagorean theorem:
 $a^2 + b^2 = c^2$
In a right-angled triangle, the square of the hypotenuse (c) is equal to the sum of the squares of the other two sides (a and b).

Euclidean (straight-line) distance:
 $d_i = \sqrt{(X_i - X_j)^2 + (Y_i - Y_j)^2}$

FIGURE 4.1
Calculation of Euclidean Distance (d_i) from Point i to Point j

TABLE 4.2							
Worktable for Calculating Central Point							
Distance matrix between points in figure 4.1							
	A	B	C	D	E	F	G
A	0	2.59	1.93	1.68	1.55	2.56	2.90
B	2.59	0	1.96	3.33	3.82	3.86	3.31
C	1.93	1.96	0	1.58	2.34	1.92	1.41
D	1.68	3.33	1.58	0	0.91	0.89	1.58
E	1.55	3.82	2.34	0.91	0	1.58	2.47
F	2.56	3.86	1.92	0.89	1.58	0	1.14
G	2.90	3.31	1.41	1.58	2.47	1.14	0

Point "D" is the lowest total distance from all other points: 9.97

Euclidean distance from point A to point B:

$$\begin{aligned}d_{AB} &= \sqrt{(X_A - X_B)^2 + (Y_A - Y_B)^2} \\&= \sqrt{(2.8 - 1.6)^2 + (1.5 - 3.8)^2} \\&= 2.59\end{aligned}$$

Euclidean (straight-line) distance

- Total distance from all other points is lowest

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Center Point

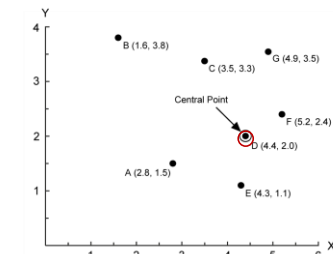


FIGURE 4.1
Graph of Locational Coordinates and Central Point

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Mean Center

- **mean center** – average location of a set of points
 - **Center of gravity** of point pattern (spatial distribution)
 - average X, Y values
 - equal weights

$$\bar{X} = \frac{\sum X_i}{n} \text{ and } \bar{Y} = \frac{\sum Y_i}{n}$$

where:

\bar{X} = mean center of X

\bar{Y} = mean center of Y

X_i = X coordinate of point i

Y_i = Y coordinate of point i

n = number of points in the distribution

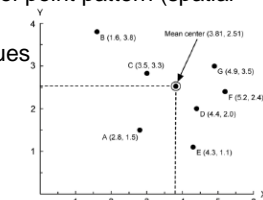


FIGURE 4.1
Graph of Locational Coordinates and Mean Center

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Mean Center

- **Outliers...**
 - add point (15, 13)
 - Average location but...

TABLE 4.2 Worktable for Calculating Mean Center		
Locational coordinates *		
Point	X_i	Y_i
A	2.8	1.5
B	1.6	3.8
C	3.5	3.3
D	4.4	2.0
E	4.3	1.1
F	5.2	2.4
G	4.9	3.5

$$n = 7 \quad \sum X_i = 29.7 \quad \sum Y_i = 17.6$$

$$\bar{X} = \frac{\sum X_i}{n} = \frac{29.7}{7} = 4.24 \quad \bar{Y} = \frac{\sum Y_i}{n} = \frac{17.6}{7} = 2.51$$

Mean center coordinates: (4.24, 2.51) *

* See Figure 4.1 for graph of locational coordinates and mean center.

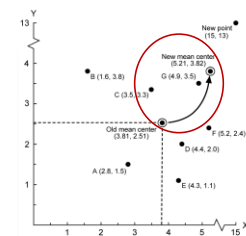


FIGURE 4.4
How an Outlier Might Affect Mean Center Location

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Mean Center

- geographic “center of population” – point where a rigid map of the country would balance if equal weights (i.e., location of each person) were situated over it

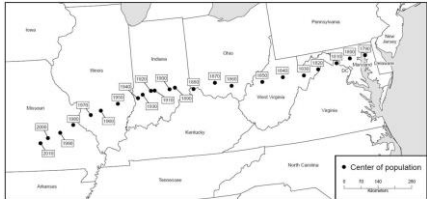


FIGURE 4.5
Geographic Center of U.S. Population, 1790 to 2010
Source: United States Bureau of the Census, 2012

Weighted mean center

- Unequal weights applied to points
 - Ex. retail store volume, city populations, etc.
 - Weights analogous to frequencies

$$\bar{X}_{wc} = \frac{\sum f_i X_i}{\sum f_i} \text{ and } \bar{Y}_{wc} = \frac{\sum f_i Y_i}{\sum f_i}$$

where \bar{X}_{wc} = weighted mean center of X
 \bar{Y}_{wc} = weighted mean center of Y
 f_i = frequency (weight) of point i

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Weighted mean center

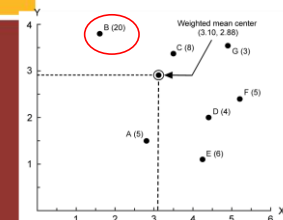


FIGURE 4.6
Graph of Point Locations, Weights (in Parentheses) and Weighted Mean Center

TABLE 4.4

Worktable for Calculating Weighted Mean Center

Point	Locational coordinates		Weight	Weighted coordinates	
	X_i	Y_i		f_i	$f_i X_i$
A	2.8	1.5	5	14.0	7.5
B	1.6	3.9	20	32.0	78.0
C	3.5	3.3	8	28.0	26.4
D	4.4	2.0	4	17.6	8.0
E	4.3	1.1	6	25.8	6.6
F	5.2	2.4	5	26.0	12.0
G	4.9	3.5	3	14.7	10.5
$n = 7$	$\sum f_i = 51$		$\sum f_i X_i = 158.1$		$\sum f_i Y_i = 147.0$
$\bar{X}_{wc} =$	$\frac{\sum f_i X_i}{\sum f_i} = \frac{158.1}{51} = 3.10$				
$\bar{Y}_{wc} =$	$\frac{\sum f_i Y_i}{\sum f_i} = \frac{147.0}{51} = 2.88$				
Weighted mean center coordinates: (3.10, 2.88)					

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Spatial measures of dispersion

- **standard distance** – measures the amount of absolute dispersion in a point distribution
 - spatial equivalent to standard deviation
 - calculate Euclidean distance from each point to mean center

$$S_D = \sqrt{\frac{\sum (X_i - \bar{X}_c)^2 + \sum (Y_i - \bar{Y}_c)^2}{n}}$$

$$S_D = \sqrt{\left(\frac{\sum X_i^2}{n} - \bar{X}_c^2\right) + \left(\frac{\sum Y_i^2}{n} - \bar{Y}_c^2\right)}$$

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Standard distance

TABLE 4.8
Worktable for Calculating Standard Distance

Point	X _i	Y _i	X _i ²	Y _i ²
A	2.0	1.0	7.04	2.25
B	1.0	3.0	2.89	16.00
C	3.0	3.0	12.25	16.00
D	4.0	2.0	16.00	4.00
E	4.0	1.0	16.00	1.00
F	3.0	2.0	9.00	4.00
G	4.0	3.0	16.00	9.00

$\bar{X}_c = 3.81$ $\bar{Y}_c = 2.51$ $\bar{X}_c^2 = 14.52$ $\bar{Y}_c^2 = 6.30$
 $n = 7$ $\sum X_i^2 = 111.90$ $\sum Y_i^2 = 50.80$
 $S_D = \sqrt{\left(\frac{111.90}{7} - 14.52\right) + \left(\frac{50.80}{7} - 6.30\right)}$
 $= \sqrt{\left(\frac{111.90}{7} - 14.52\right) + \left(\frac{50.80}{7} - 6.30\right)}$
 $= 1.54$

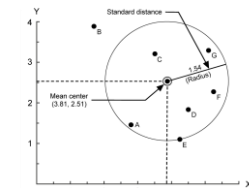


FIGURE 4.10
Graph of Point Locations, Mean Center, and Standard Distance

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Relative Measure...

Weighted standard distance

- Used with weighted mean center

$$S_{WD} = \sqrt{\left(\frac{\sum f_i(X_i^2)}{\sum f_i} - \bar{X}_{wc}^2\right) + \left(\frac{\sum f_i(Y_i^2)}{\sum f_i} - \bar{Y}_{wc}^2\right)}$$

$$S_{WD} = \sqrt{\left(\frac{\sum f_i(X_i^2)}{\sum f_i} - \bar{X}_{wc}^2\right) + \left(\frac{\sum f_i(Y_i^2)}{\sum f_i} - \bar{Y}_{wc}^2\right)}$$

- Difference
 - 1.54 vs. 1.70

TABLE 4.9
Worktable for Calculating Weighted Standard Distance

Point	f _i	X _i	X _i ²	f _i X _i ²	Y _i	Y _i ²	f _i Y _i ²
A	5	2.0	7.04	35.20	1.5	2.25	11.25
B	20	1.0	2.89	57.80	3.0	9.00	180.00
C	9	3.0	12.25	110.25	3.0	9.00	81.00
D	4	4.0	16.00	64.00	2.0	4.00	16.00
E	6	4.0	16.00	96.00	1.0	1.00	6.00
F	8	3.0	9.00	72.00	2.0	4.00	32.00
G	3	4.0	16.00	48.00	3.0	9.00	27.00

From earlier calculation of weighted mean center:

$$\bar{X}_{wc} = 3.10 \quad \bar{Y}_{wc} = 2.88 \quad \bar{X}_{wc}^2 = 9.61 \quad \bar{Y}_{wc}^2 = 8.29$$

$$\sum f_i = 51 \quad \sum f_i(X_i)^2 = 594.01 \quad \sum f_i(Y_i)^2 = 475.98$$

$$S_{WD} = \sqrt{\left(\frac{\sum f_i(X_i^2)}{\sum f_i} - \bar{X}_{wc}^2\right) + \left(\frac{\sum f_i(Y_i^2)}{\sum f_i} - \bar{Y}_{wc}^2\right)}$$

$$= \sqrt{\left(\frac{594.01}{51} - 9.61\right) + \left(\frac{475.98}{51} - 8.29\right)}$$

$$= 1.70$$

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Standard Deviational Ellipse

- Extends standard distance to include orientation of the point pattern
 - Calculated separately for X and Y
 - Average distance points vary from mean center on X and average distance points vary

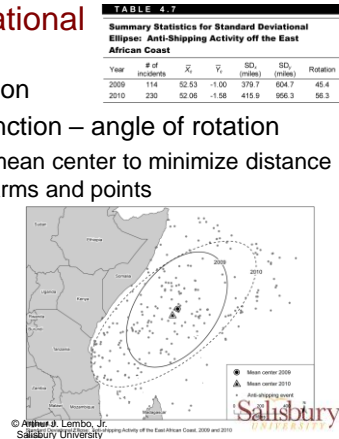
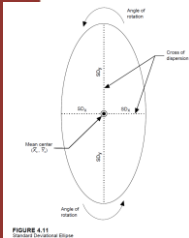
$$SD_x = \sqrt{\frac{\sum (X_i - \bar{X})^2}{n}} \text{ and } SD_y = \sqrt{\frac{\sum (Y_i - \bar{Y})^2}{n}} \quad (4.16)$$

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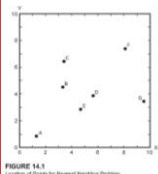
Standard Deviational Ellipse

- cross of dispersion
- trigonometric function – angle of rotation
 - Rotated about mean center to minimize distance between both arms and points



Nearest Neighbor Analysis – (NNA)

- Distance of each point to its nearest neighbor is measured and mean distance for all points is determined
 - Objective: describe the pattern of points in a study region and make inferences about the underlying process



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Nearest Neighbor analysis – (NNA)

- Compare calculated value from point data to theoretical point distributions
 - Outcomes: random, clustered, dispersed
 - average nearest neighbor distance is an absolute index
 - Dependent on distance measure (ex. miles, km, meters, etc.)
 - Minimum = 0 (clustered), maximum is function of point density
 - standardized nearest neighbor index (R) is often used
 - Comparison of data to random

$$R = \frac{\overline{NND}}{NND_R} \quad (14.5)$$

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Nearest Neighbor analysis – (NNA)

$$\overline{NND} = \frac{\sum NND}{n} \quad (14.1)$$

where n = number of points.

$$\overline{NND}_R = \frac{1}{2\sqrt{Density}} \quad (14.2)$$

where \overline{NND}_R = mean nearest neighbor distance in a random pattern

Density = number of points (n) / Area

$$\overline{NND}_R = \frac{1}{2\sqrt{0.7}} = 1.89$$

$$R = \frac{\overline{NND}}{\overline{NND}_R} \quad (14.5)$$

TABLE 14.1
Coordinates and Nearest Neighbor Information for Example*

Point	X	Y	NV	NND
A	1.3	9.9	E	3.94
B	3.2	4.4	C	2.80
C	3.3	6.4	B	2.00
D	5.6	3.6	E	1.38
E	4.8	2.7	D	1.36
F	8.1	7.4	G	4.21
G	9.4	3.4	D	3.82
SUM				18.69

where NV = nearest neighbor
NND = nearest neighbor distance

$$\overline{NND} = \frac{\sum NND}{n} = \frac{18.69}{7} = 2.67$$

* See Figure 14.1 for graph of point locations.

$$R = \frac{2.67}{1.89} = 1.41$$

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NNA – R values

- Continuum...
 - Result?
 - Descriptive test.

$$R = \frac{2.67}{1.89} = 1.41$$

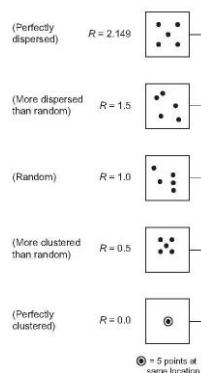
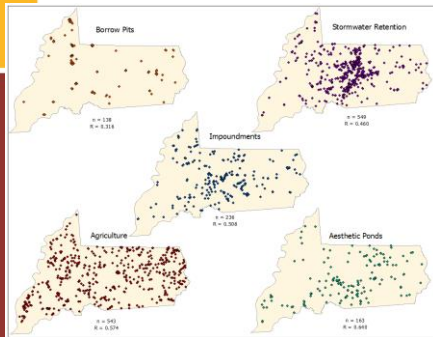


FIGURE 14.2
Continuum of R Values in Nearest Neighbor Analysis
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Functional
SWB

Results?

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Nearest neighbor analysis (nna)

- A difference test can be used to determine if the observed nearest neighbor index (NNA) differs significantly from the theoretical norm (NNA_R)
 - H_0 : There is no difference between our distribution and a random distribution (Poisson)

$$Z_s = \frac{NND - NND_s}{\sigma_{NND}} \quad (14.6)$$

where σ_{NND} = standard error of the mean nearest neighbor distances

The standard error for the nearest neighbor test can be estimated with the following formula:

$$\sigma_{NND} = \frac{.26136}{\sqrt{n(Density)}} \quad (14.7)$$

where: n = number of points $Density$ = number of points (n)/Area

Nearest Neighbor Analysis	
Primary Objective:	Determine whether a random (Poisson) process has generated a point pattern
Requirements and Assumptions:	1. Random sample of points from a population 2. Sample points are independently selected
Hypotheses:	$H_0: NND = NND_s$ (point pattern is random) $H_1: NND \neq NND_s$ (point pattern is not random) $H_2: NND > NND_s$ (point pattern is more dispersed than random) $H_3: NND < NND_s$ (point pattern is more clustered than random)
Test Statistic:	$Z_s = \frac{NND - NND_s}{\sigma_{NND}}$

Nearest neighbor analysis (nna) Example: Community Services in Toronto

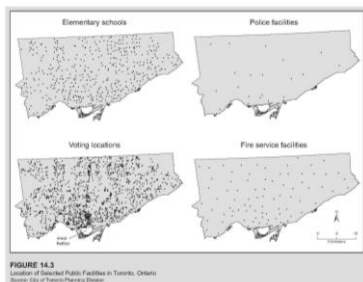


FIGURE 14.3
Location of Selected Public Facilities in Toronto, Ontario
Source: City of Toronto Planning Division

- Emergency services: fire and police
 - Seek dispersion to provide services equally
- Nonemergency services: polling sites and elementary schools
 - Seek clustering...why?

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Nearest neighbor analysis (nna) Example: Community Services in Toronto

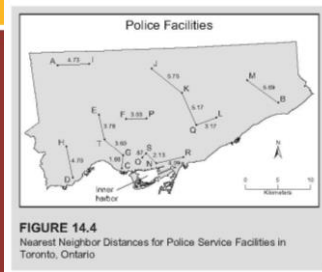


FIGURE 14.4
Nearest Neighbor Distances for Police Service Facilities in Toronto, Ontario

- Result?

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TABLE 14.3
Worktable for Nearest Neighbor Analysis: Police Service Facilities in Toronto, Ontario

$H_0: XND = XND_0$ (point pattern is random)
 $H_a: XND < XND_0$ (point pattern is not random)

Calculate mean nearest neighbor distance:

$$XND = \frac{\sum XND_i}{n} = \frac{71.65}{20} = 3.58$$

where: n = number of points

Calculate random nearest neighbor distance:

$$XND_0 = \frac{1}{2} \sqrt{\frac{4A}{n}}$$

where: Density = n/A

$$XND_0 = \frac{1}{2} \sqrt{\frac{4(706)}{20}} = \frac{1}{2} \sqrt{141.2} = 2.53$$

Calculate standardized nearest neighbor distance:

$$Z = \frac{XND - XND_0}{\frac{XND_0}{2.25}} = \frac{3.58 - 2.53}{\frac{2.53}{2.25}} = 1.20$$

Calculate test statistic:

$$Z = \frac{XND - XND_0}{\frac{XND_0}{2.25}}$$

where: Density = n/A

$$Z = \frac{3.58 - 2.53}{\frac{2.53}{2.25}} = \frac{1.05}{1.14} = 0.92$$

$Z = \frac{1.05 - 0.00}{0.92} = 1.15$ (p = 0.0675)

Nearest neighbor analysis (nna) Example: Community Services in Toronto

TABLE 14.3
Nearest Neighbor Values for Selected Public Facilities in Toronto, Ontario

Public Facility	NND	Density	NND ₀	Z	p
Fire	2.26	1.278	1.42	0.21	.0000
Police	3.03	0.312	2.03	1.28	.0103
Elementary schools	0.59	7.358	0.58	0.95	.3430
Voting locations	0.27	2.4351	0.32	-12.23	.0000

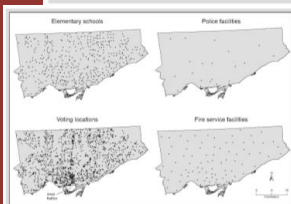
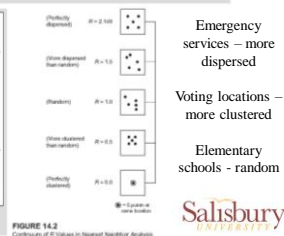


FIGURE 14.3
Location of Selected Public Facilities in Toronto, Ontario



Nearest neighbor analysis (nna) Example: Community Services in Toronto

- Issues to consider...
 - Study area boundaries – political boundary or research delimited
 - Doesn't impact NNA distances but does impact area (point density function)
 - Nearest feature – may be outside study area!
 - Problem with using political boundaries
 - More advanced techniques available – Ripley's K
 - Evaluates more than one nearest neighbor
 - Can define distances – How many police stations within 1km? 2km?

$$\sigma_{NND} = \frac{.26136}{\sqrt{n(Density)}} \quad (14.7)$$

where: n = number of points
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General Issues in Inferential Spatial Statistics

- Geographers are interested in spatial patterns produced by physical or cultural processes
 - Explain patterns of points and areas
 - "global" overall arrangement
 - Random vs. Nonrandom spatial processes
 - "local" concentrations or absences
 - Clusters – points or areas within larger area
 - Groups of high values – "hot spots"
 - Groups of low values – "low spots"

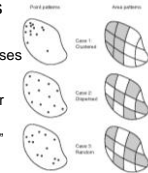


FIGURE 13.1
Types of Point and Area Patterns

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Types of Spatial Patterns

- Compare existing pattern to theoretical pattern

Clustered

- Density of points varies significantly from one part of study area to another
 - Points: retail locations near highway interchange
 - Areas: registered majority political party affiliation
- Patterns result from nonrandom factors
 - Accessibility, income, race, etc.

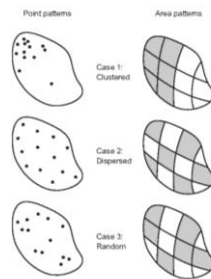


FIGURE 13.1
Types of Point and Area Patterns

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Types of Spatial Patterns

Dispersed

- Uniformly distributed across study area
 - Suggests systematic spatial process
- Area example: Central Place Theory
 - settlements are uniformly distributed across landscape to best serve needs of a dispersed rural population

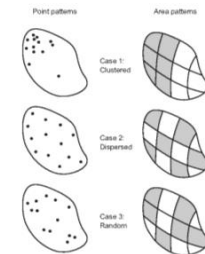


FIGURE 13.1
Types of Point and Area Patterns

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Types of Spatial Patterns

- **Random**
 - No dominant trend toward clustering or dispersion
 - Suggests spatially random process (Poisson)
 - Ex. lightning strikes
- **Geographic problems**
 - Patterns typically appear as some combination of these three patterns
 - Along continuum...

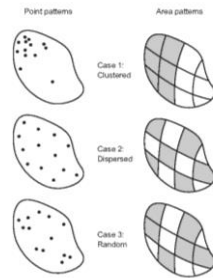


FIGURE 13.1
Types of Point and Area Patterns

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Spatial Autocorrelation

- Tobler's Law – "Everything is related to everything else but near things are more related than distant things"
- **spatial autocorrelation**: measures the degree to which a geographic variable is correlated with itself through space
 - Positive, negative or non-existent
 - Positive spatial autocorrelation: objects near one another tend to be similar
 - Features with high values are near other features with high values, features with medium values are near other features with medium values, etc.
 - Negative spatial autocorrelation: objects near one another tend to have sharply contrasting values
 - Features with high values near features with low values
- Most geographic phenomena exhibit positive spatial autocorrelation
 - Examples: rainfall amounts, home values, etc.

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Variogram

- Visualization of spatial autocorrelation
- **variogram**: scatterplot that display the differences in values between geographic locations against the differences in distances between the geographic locations
 - Y-axis: average **variance** (really half the variance) in **values** for a set of geographic objects
 - X-axis: **distance** between objects
 - Use plot to determine average difference in values at specific distances
 - Ex. 100 miles, 500 miles

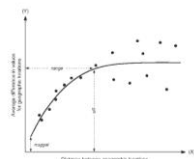


FIGURE 13.2
An Example of a Variogram

Geographic locations near one another tend to have smaller differences than geographic locations at greater distances (positive autocorrelation)!

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Variogram

- Displayed as best-fitting curve (function)
 - Differences in values with distance noted and then diminishes
- **range** - distance at which the difference in values are no longer correlated
- **sill** - average difference in value where there is no relationship between location and value
- **nugget** - degree of uncertainty when measuring values for geographic locations that are very close to one another
 - Effect of sampling, measurement error, etc.
 - Unlikely that two samples near each other will have the exact same value

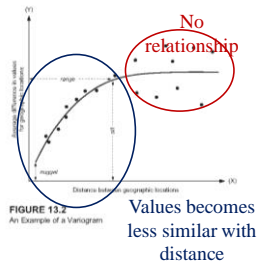
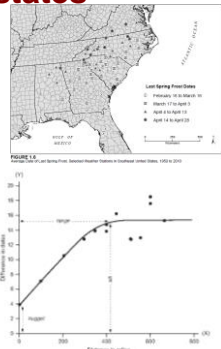


FIGURE 13.2
An Example of a Variogram

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Variogram Example: Last Spring Frost IN SE United states

- Two nearby stations, LSF dates should be similar
 - 0 to 400 miles: distances between stations are large, dates are different
 - Beyond 400 miles, no longer spatially autocorrelated...



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FIGURE 13.3
Variogram of Last Spring Frost Dates in Southeastern United States, 1980 to 2010

Spatial Autocorrelation: Importance in Geographic Research

- GIS - push of a button
 - Calculates relationship for any distances...
 - Is the test appropriate for any distance?
- Presence of spatial autocorrelation
 - Inferential statistics assume independent observations
 - Example: last spring frost dates are spatially correlated!
 - Impact: sample locations close together, just like taking the same sample
 - Sample size impacts size of standard error
 - Smaller standard error than warranted
 - Standard deviation calculation impacted
 - Even smaller standard error
- Global or local measurement
 - **global** - examine a distribution of subset (ex. ethnic group) across entire area (ex. city)
 - One group more clustered, dispersed or random than another
 - **local** - compares each geographic object (ex. all group members) with its surrounding neighbors
 - Is area (ex. neighborhood) more clustered, dispersed or random than another?

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Spatial Autocorrelation: Neighbor Definitions

- Measure of interaction between geographic features
 - Defining neighbor...
 - **adjacency** – share common border
 - Binary: yes or no
 - » Ex. New York and Pennsylvania, New York and California
 - **distance threshold** – cut-off distance
 - Salisbury, MD – neighbor definition 60 miles...Easton, Wilmington, DE?
 - **inverse-distance** – strength of “neighborliness” between two objects as a function of distance separating them ($1/\text{distance}$)
 - » New York City and Boston: 1/189 miles or .005,
 - » NYC and LA: 1/2588 miles or .0004
 - » Interaction measure (“neighborliness”) is 12 times stronger between NYC and Boston versus NYC and LA
 - In equations/modeling, takes the form of weights
 - w_{ij} : weight between geographic object i and j
 - Binary: 0 or 1
 - Inverse-distance: continuous value ...

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Spatial Autocorrelation

- First law of geography: “everything is related to everything else, but near things are more related than distant things” – Waldo Tobler
- Many geographers would say “I don’t understand spatial autocorrelation” Actually, they don’t understand the mechanics, they do understand the concept.

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Spatial Autocorrelation

- Spatial Autocorrelation – correlation of a variable with itself through space.
 - If there is any systematic pattern in the spatial distribution of a variable, it is said to be spatially autocorrelated
 - If nearby or neighboring areas are more alike, this is *positive spatial autocorrelation*
 - *Negative autocorrelation* describes patterns in which neighboring areas are unlike
 - Random patterns exhibit *no spatial autocorrelation*

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Why spatial autocorrelation is important

- Most statistics are based on the assumption that the values of observations in each sample are independent of one another
- Positive spatial autocorrelation may violate this, if the samples were taken from nearby areas
- Goals of spatial autocorrelation
 - Measure the strength of spatial autocorrelation in a map
 - test the assumption of independence or randomness

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Spatial Autocorrelation

- Spatial Autocorrelation is, conceptually as well as empirically, the two-dimensional equivalent of redundancy
- It measures the extent to which the occurrence of an event in an areal unit constrains, or makes more probable, the occurrence of an event in a neighboring areal unit.

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Spatial Autocorrelation

- Non-spatial independence suggests many statistical tools and inferences are inappropriate.
 - Correlation coefficients or ordinary least squares regressions (OLS) to predict a consequence assumes that the observations have been selected randomly.
 - If the observations, however, are spatially clustered in some way, the estimates obtained from the correlation coefficient or OLS estimator will be biased and overly precise.
 - They are biased because the areas with higher concentration of events will have a greater impact on the model estimate and they will overestimate precision because, since events tend to be concentrated, there are actually fewer number of independent observations than are being assumed.

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Indices of Spatial Autocorrelation

- Moran's I
- Geary's C
- Ripley's K

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Moran's I Index (Global)

- Popular technique for quantifying level of spatial autocorrelation in a set of geographic areas
- Moran's I Index takes into account geographic locations (points or areas) as well as attribute values (ordinal or interval/ratio) to determine if areas are clustered, randomly located or dispersed
 - Positive : clustered – nearby locations have similar attribute values
 - Negative: dispersed – nearby locations have dissimilar attribute values
 - Near zero: attribute values are randomly dispersed throughout study area

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Moran's I Index (Global)

The general form of Moran's Index for areas is shown in equation 15.4, and the mathematical equivalent, as presented in Ebdon (1988), appears in equation 15.5:

$$I = \frac{\left(\frac{\text{number of areas}}{\text{number of joins}} \right) \left(\frac{\text{sum of cross-products for all contiguous pairs } (i,j)}{\text{variance of the area attribute values}} \right)}{\quad} \quad (15.4)$$

$$I = \frac{n \sum (x_i - \bar{x})(x_j - \bar{x})}{f \sum (x - \bar{x})^2} \quad (15.5)$$

where:

n = the number of areas
 x = an area attribute value
 \bar{x} = the mean of all area attribute values
 x_i and x_j = the values of contiguous pairs

$\sum (x_i - \bar{x})(x_j - \bar{x})$ = the sum of all contiguous pairs
 f = the number of joins
 $\sum (x - \bar{x})^2$ = the variance of the attribute values.

Weighted cross-products: deviation values for contiguous pairs multiplied together and summed

- Positive: neighboring areas with similar attribute values either large or small (clustered)
- Larger deviation from mean, greater magnitude
- Negative: neighboring areas with dissimilar attribute values contiguous (dispersed)
- Larger deviation from mean, greater magnitude
- Near zero: random...

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• I ranges from -1.00 to 1.00

Moran's I Index (Global): Significance test

The null hypothesis is that the area values are arranged in a completely random spatial pattern, and that spatial autocorrelation is not present within the study area. The Z score is computed as:

$$Z = \frac{I - E_I}{\sqrt{\text{Var}_I}} \quad (15.6)$$

where: $E_I = \frac{-1}{(n-1)}$

Similar to the free and non-free sampling techniques mentioned with joint count analysis and binary data, the variance for Moran's I may be calculated under the assumption of normality or randomization. Normality is similar to free sampling in that the observed values of a variable are taken randomly from a normally distributed population, whereas randomization is similar to non-free sampling in that the values are taken directly from the data set. The variance under the normality assumption is computed as:

$$\text{Var}_I = \frac{n^2 J + 3J^2 - n \sum L_i^2}{J^2 (n^2 - 1)}$$

where: n = the number of objects
 J = the number of joins, and
 $\sum L_i$ = the sum of the number of joins for each individual area

H_0 : No spatial autocorrelation in the data (Values of areas are completely random)

H_A : Spatial autocorrelation in the data (Values of areas are not completely random)

- If p-value is not significant, then you should not reject the null hypothesis

- The observed pattern is not different from complete spatial randomness

- p-value significant and Z-score

Variance under the randomization assumption is computed as:

$$\text{Var}_I = \frac{n^2 [n^2 + 2 - 3n] - 3J^2 - n \sum L_i^2 - 3J [n^2 - n] - 6J^2 - 2n \sum L_i^2}{J^2 [n^2 - 3n - 3] - 2n \sum L_i^2}$$

where: L_i represents the histogram for the variable x

itive
 stering
 ant and Z-score
 tive

TABLE 13.6
Worktable for Moran's I: Five-area Example Pattern

H_0 : no spatial autocorrelation in the data
 H_A : spatial autocorrelation exists

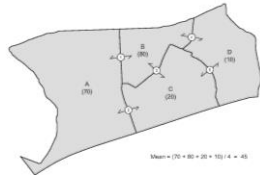
Area	1	2	3	4	5
A	2	4	1	1	1
B	3	3	1	1	1
C	3	3	1	1	1
D	2	4	1	1	1
E	2	4	1	1	1

Join	1	2	3	4	5
A	2	4	1	1	1
B	3	3	1	1	1
C	3	3	1	1	1
D	2	4	1	1	1
E	2	4	1	1	1

Join number	x_i	$(x_i - \bar{x})$	x_j	$(x_j - \bar{x})$	$(x_i - \bar{x})(x_j - \bar{x})$
1	70	25	80	35	875
2	70	25	20	-25	-625
3	20	-25	80	35	-875
4	80	35	10	-35	-1225
5	10	-35	20	-25	875
SUM					-975

$J = 5$
 $\sum x_i^2 = 26$
 $\bar{x} = \frac{26}{5} = 5.2$
 $\sum (x_i - \bar{x})^2 = 3,700$

$I = \frac{-1}{(n-1)} = \frac{-1}{4} = -0.25$
 $I = \frac{\sum (x_i - \bar{x})(x_j - \bar{x})}{J \sum (x_i - \bar{x})^2} = \frac{-975}{5 \times 3,700} = -0.21$
 $\text{Var}_I = \frac{J^2 + 3J^2 - n \sum L_i^2}{J^2 (n^2 - 1)} = \frac{(5^2 + 3 \times 5^2 - 5 \times 26)}{5^2 (5^2 - 1)} = \frac{51}{375} = 0.136$
 $Z = \frac{I - E_I}{\sqrt{\text{Var}_I}} = \frac{-0.21 - (-0.25)}{\sqrt{0.136}} = 0.326$
 $p = 0.74$



Result?

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Example: Cleveland Census Block Groups

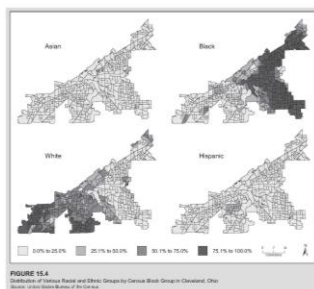
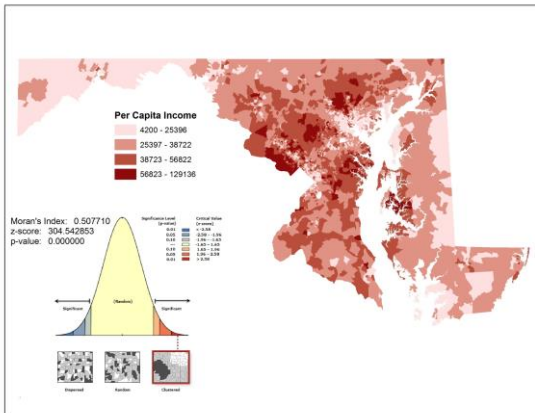
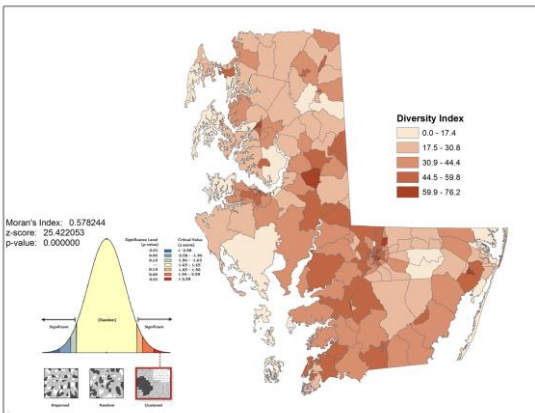


TABLE 13.7
Moran's I Values and Test Statistics for Selected Racial and Ethnic Groups in Cleveland, Ohio Census Block Groups

Group	Moran's I	Z	p-value
Asian	0.35	16.28	0.000
Black	0.58	64.6	0.000
White	0.64	65.4	0.000
Hispanic	0.81	65.7	0.000

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Moran's I Index (Global)

Moran's Index

Primary Objective: Identify significant spatial patterns within a study area

Requirements and Assumptions:

- Minimum of (30) geographic features
- Attribute values measured on an ordinal or interval/ratio scale

Hypotheses:

H_0 : Attribute values are randomly distributed across features in the study area

H_a : Attribute values are not randomly distributed across features in the study area

Test Statistic:

$$I = \frac{n \sum (x_i - \bar{x})(x_j - \bar{x})}{\sum \sum (x_i - \bar{x})^2}$$

Interpretation:

Assuming a significant p-value:

- $I < 0$ (observed pattern is dispersed)
- $I = 0$ (observed pattern is random)
- $I > 0$ (observed pattern is clustered)

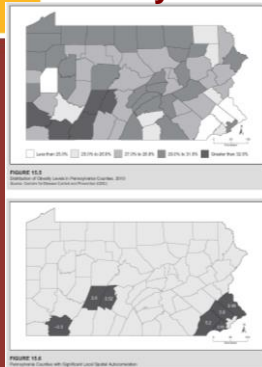
Moran's I Index (local)

- Global spatial autocorrelation (Moran's I) may indicate a lack of spatial autocorrelation
 - Local pockets may exist– hotspots
 - LISA – Local Indicators of Spatial Association
 - Quantify similarity of each geographic observation with an identified group of geographic neighbors
 - Identifies local clusters – geographic locations where adjacent or nearby areas have similar values
 - Spatial outliers – geographic locations that are different from adjacent or nearby areas
 - Each geographic area receives individual measure

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Moran's I Index (local): Example: Obesity in PA



$$I = \frac{(x_i - \bar{X})}{S_i^2} \sum_{j \in N(i)} w_{ij} (x_j - \bar{X}) \quad (15.7)$$

$$S_i^2 = \sum_{j \in N(i)} \frac{(x_j - \bar{X})^2}{n-1} - \bar{X}^2$$

where: x_i = the value for a particular geographic entity
 x_j = the value for the neighboring geographic entity
 \bar{X} = the average of all attributes
 w_{ij} = the spatial weights.

The overall computation of the local Moran index is beyond the scope of this book, but a short example will help illustrate its usefulness.

- Global Moran's I = .69, p-value = .25
- Local Moran's I for each county...

Positive values: similar levels in adjacent counties (clustering)

- Philly...
- Johnstown/Altoona

Negative values: dissimilar values – outlier

- Fayette County

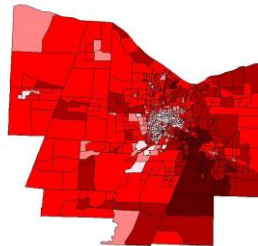
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Example of Moran's I – Per Capita Income in Monroe County

Using Polygons:
Moran's I: .66
P: < .001

Using Points:
I: .12
Z: 65



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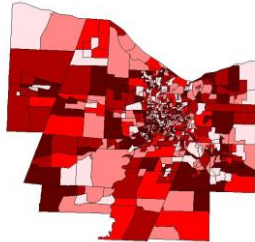
Example of Moran's I – Random Variable

Using Polygons:

Moran's I: .012
p: .515

Using Points:

Moran's I: .0091
Z: 1.36



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